

1 **Economic valuation of ecosystem services provided by irrigation of**
2 **meadows and pastures: evidence from Aosta Valley, Italy**

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19
20 **Abstract**

21 *This study examines affordability issues related to EU water pricing policies in mountain areas, where*
22 *irrigation entails costs but also provides valuable ecosystem services. Focusing on Italy's Aosta Valley,*
23 *we use a discrete choice experiment to assess residents' willingness to pay (WTP) for ecosystem services*
24 *from irrigated meadows and pastures, specifically: maintaining active livestock farms, reducing*
25 *landslide risk, attracting summer tourism, and preserving the traditional landscape. Respondents valued*

26 *the traditional landscape at €126/year, each livestock farm at €0.5/year, and each landslide avoided at*
27 *€3.4/year; summer tourism was not statistically significant. Total WTP estimated is €633.5/year per*
28 *household, or approximately €36 million/year for the region. Results support recognising traditional*
29 *mountain irrigation as a provider of ecosystem services and inform the integration of their value into*
30 *EU water policy.*

31

32 **Keywords:** ecosystem services; irrigation; livestock farms; mountain areas; Discrete Choice Experiment

33 **JEL Codes:** Q25, Q57; H41

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37 1. INTRODUCTION

38 Facing growing pressure on water resources, the European Union (EU) has promoted more sustainable
39 water management through the Water Framework Directive 2000/60 (WFD), which aims to establish
40 overarching principles for sustainable water management across sectors. In agriculture, this approach is
41 primarily economic and based on the “polluter pays principle”. According to this principle, those who
42 pollute or exploit natural resources are obliged to cover the associated costs, incentivizing more efficient
43 and conscious use (Buttinelli et al., 2024; Zucaro et al., 2015). To this end, Article 9 of the WFD
44 introduced a water-pricing policy for irrigation water, aimed to “provide an incentive for users to use
45 water resources efficiently, and thereby contribute to the environmental objectives of the Directive”. The
46 European Commission (2000) specified that users should recover three cost components: financial costs,
47 environmental costs, and resource costs.

48 Estimating these components has proved difficult across Member States, especially in marginal
49 agricultural areas where higher water charges raise affordability concerns (Zucaro et al., 2015).

50 In Italy, irrigation water services are managed by regional administrations and implemented by Water
51 User Associations (WUAs). Implementing the Directive has been particularly challenging in areas with
52 environmental and socio-economic constraints, such as mountain regions. In the Aosta Valley, the EU
53 has criticised the Region for the absence of a coherent irrigation water pricing system. Here, deeply
54 rooted traditional practices, further complicate the picture: WUA members contribute voluntary labour
55 to the regular maintenance of irrigation networks, effectively constituting a non-monetary form of cost
56 recovery not recognised as compliant under EU guidelines.

57 The situation was further compounded when the EU made implementation of the WFD pricing policy a
58 condition for accessing European funds, including Rural Development funds (Regulation (EU)
59 1303/2013). In response, the regional administration argues that irrigation water use for mountain
60 meadows and pastures entails not only costs, but also external benefits in the form of public goods and
61 services.

62 Particularly in inland and marginal areas, dairy livestock farming is a vital activity that provides and
63 maintains various ecosystem services (ES), including biodiversity, water flow regulation, landscape,
64 recreation, and cultural benefits (Pauler et al., 2025; Herzog & Seidl, 2018; Orlandi et al., 2016;
65 Battaglini et al., 2014). To implement a comprehensive compensation criterion, the value of these ES
66 should be estimated. Assigning a monetary value to these services allows them to be recognised at the
67 European level and to be compared with the recovery costs established by the water pricing policy.

68 A growing body of literature has explored the social preferences for agricultural ES across European
69 countries (Alcon et al., 2020; Bernués et al., 2019; Tienhaara et al., 2019; Novikova et al., 2017;
70 Chatzinikolaou et al., 2015).

71 Many studies have also focused on the economic valuation of irrigation-related ES, highlighting their
72 role in promoting societal well-being and estimating willingness to pay (WTP) for improvements in
73 water quality, drought mitigation and landscape conservation (Khan et al., 2019; Vivithkeyoonvong,
74 2017). These services are increasingly recognised for their non-market value (Kourtis & Tsihrintzis,
75 2017). Stated preference studies reveal that people show WTP for the protection and restoration of water
76 ecosystems, especially where agricultural practices or infrastructure improvements can enhance local
77 ecological conditions (Kourtis & Tsihrintzis, 2017; Lee et al., 2017). These ES also generate broader
78 societal benefits, including biodiversity conservation, climate regulation, and cultural heritage
79 preservation (Khan et al., 2019). Overall, integrating their value into policy could better balance
80 environmental protection with agricultural and water management practices, supporting more
81 sustainable outcomes (Khan et al., 2019; Vivithkeyoonvong, 2017).

82 The value of agricultural ES, including water quality, biodiversity, landscape preservation and
83 hydrogeological regulation, has also been explored in northern Italian regions (Blasi et al., 2023;
84 Mazzocchi & Sali, 2022; Tempesta & Vecchiato, 2018; Faccioni et al., 2019), where mountain
85 agricultural practices help sustain fragile ecosystems and provide public goods beyond economic output,
86 and where citizens have shown a clear preference for agricultural practices that prioritise environmental
87 services over intensification.

88 In particular, Ruberto et al. (2022) analysed traditional irrigation systems in Lombardy, highlighting
89 their role in providing ES such as water regulation and biodiversity conservation, and found a significant
90 WTP for higher levels of these services, suggesting that their inclusion in water pricing policies under
91 the WFD could encourage more sustainable agricultural water use. Similarly, Tempesta and Vecchiato
92 (2013) assessed residents' preferences for policies aimed at reducing nitrate concentrations in
93 groundwater and improving riverscape quality along the Serio River in northern Italy, finding that
94 residents valued both groundwater quality, due to its health implications, and broader riverscape
95 improvements linked to recreation and biodiversity. More broadly, the literature stresses the need to
96 better recognise ES within EU sectoral policies, as current instruments still struggle to integrate their
97 value into effective territorial strategies (Winkler et al., 2025; Viaggi, 2018).

98 Although research on agricultural ES is expanding, the valuation of services provided by water use in
99 marginal and mountainous areas remains poorly explored. This study addresses this gap through a
100 context-specific analysis of the ES provided by traditional irrigation practices in the extensive mountain
101 pastures of Aosta Valley, estimating regional residents' WTP as direct beneficiaries of these services,
102 and contributing to more informed water management policies that account for both the environmental
103 benefits and the socio-economic challenges of mountain areas.

104 To this end, the estimation relied on a Discrete Choice Experiment (DCE), a widely-used method in
105 environmental economics that allows for the monetary evaluation of non-market goods by presenting
106 respondents with hypothetical scenarios involving choices among alternative policies (Aizaki, 2010;
107 Alriksson & Öberg 2008; Holmes et al, 2017).

108

109 **2. SURVEY DESIGN AND DATA**

110 *2.1 Experiment setting*

111 DCEs simulate real-life decision-making by presenting respondents with multiple choice sets, each
112 combining different attribute levels of the good or scenario being valued. By including a monetary

113 attribute, it is possible to elicit WTP for different attribute level combinations (Louviere, 2001; Louviere
114 et al., 2010). The method is grounded in the Lancasterian utility theory (Lancaster, 1966) and the
115 Random Utility Theory (McFadden, 2001).

116 In analytical term, the utility that individual i derives from alternative j is expressed as:

$$117 \quad U_{ij} = V_{ij} + \varepsilon_{ij}$$

118 where V_{ij} denotes the systematic component of utility, typically specified as a function of the attributes
119 and their associated parameters, and ε_{ij} represents the stochastic component (McFadden, 2001). As a
120 consequence of this unobservable component, utility cannot be directly observed, and choices are
121 therefore modelled probabilistically. Individuals are assumed to select the alternative that yields the
122 highest utility (Revelt & Train, 1998).

123 Our experiment was designed to assess the value of ES provided by irrigation of meadows and pastures
124 in the Aosta Valley. A detailed description of the study area is provided in the Supplementary Materials.

125 Respondents were presented with a hypothetical scenario in which the regional administration
126 implements a water pricing policy in line with the WFD. As a result, local farmers may not be able to
127 afford the higher irrigation costs, with negative effects on production and the loss of some irrigation ES.
128 To limit such loss of benefits, respondents were asked to support farmers by contributing to their
129 irrigation water costs. The ES considered in the analysis were those identified in a study conducted in
130 the same area by Novelli et al. (2022).

131 Focus group discussions have been used to identify and prioritise the benefits that irrigation provides in
132 the Region. Local stakeholders identified four main benefits: i) maintaining the economic and social
133 vitality of mountain areas through the presence of livestock, ii) preventing hydrogeological instability
134 by maintaining irrigation networks, iii) attracting summer tourists for hiking and excursions in a well-
135 maintained environment, and iv) preserving the traditional agricultural landscape.

136 Following the ES framework adopted in a DCE by Faccioni et al. (2019) and consistent with established
137 classification systems (Millennium Ecosystem Assessment, 2005; Haines-Young, 2023), the benefits
138 considered in this study belong to different ES categories and contribute to distinct components of the
139 Total Economic Value (TEV). The presence of livestock farms represents a socio-economic benefit
140 derived from agroecosystem functioning. Cultivated crops (including those providing fodder for
141 livestock) and animals reared for nutritional purposes are classified as *provisioning ecosystem services*
142 and correspond to a consumptive direct use value (Haines-Young, 2023; FAO, 2025). The prevention of
143 hydrogeological instability constitutes a *regulating ecosystem service*, generating an indirect use value
144 by reducing environmental risks and safeguarding public safety and infrastructure. Landscape and
145 tourism are a *cultural ecosystem services* that provide non-consumptive direct use value, reflecting their
146 aesthetic and recreational benefits. Although mass tourism in fragile areas can have negative
147 environmental impacts (Pueyo-Ros, 2018), stakeholders emphasised that summer tourism in the Aosta
148 Valley largely takes the form of nature-based ecotourism and can therefore be regarded as a cultural
149 benefit (Viddaller & Dutoit, 2022).

150 Although some ES in the DCE, particularly provisioning services and tourism-related activities, are
151 linked to market revenues, the experiment does not estimate their market value. Instead, it assesses the
152 marginal economic value and welfare effects of changes in ES provision, expressed through individuals'
153 WTP.

154 Therefore, four attributes were identified: i) number of active livestock farms, ii) landslides and debris
155 flows avoided/year, iii) summer tourists/year, and iv) landscape appearance - each measuring changes
156 in the provision level of the selected ES. Respondents were presented with a future scenario in which
157 WFD implementation would threaten the supply of these benefits over 30 years. Drawing on an approach
158 used for endangered species valuation, where respondents are asked to value the avoidance of further
159 decline (Bartczak & Meyerhoff, 2013; Wallmo & Lew, 2016; Imamura et al., 2020), participants were
160 asked to state their WTP to prevent ES deterioration. In our experimental design, the *status quo* reflects
161 the current situation with fully operational irrigation systems and maximum ES provision for all non-


162 cost attributes. The non-intervention scenario assumes that, in the absence of financial support to
163 farmers, WFD-induced cost increases progressively reduce irrigation activities, leading to ES
164 deterioration over the 30-year horizon and corresponding to the worst projected levels of non-cost
165 attributes.

166 For each attribute, three levels were defined: (i) the current situation, corresponding to the full
167 maintenance of ES through a complete offset of the additional irrigation costs arising from the
168 implementation of the WFD pricing policy; (ii) an intermediate level, reflecting a partial offset of these
169 additional irrigation costs and, consequently, a partial mitigation of the associated negative impacts; and
170 (iii) a non-intervention scenario, representing the maximum projected deterioration of ES in the absence
171 of financial support. All attributes were quantitative, except for the landscape. Landscape appearance
172 was depicted with photos showing three alternative attribute levels based on irrigation practices: green
173 meadows with cows grazing (no-change scenario), yellowed pastures, and yellowed pastures with
174 invasive tree species (non-intervention scenario). While the ES included in the DCE are interrelated
175 through irrigation, attribute levels were varied independently in the experimental design to enable the
176 estimation of marginal utilities, with each level reflecting realistic and plausible changes in the provision
177 of the corresponding service (Bateman, 2002).

178 The payment vehicle and the cost attribute levels (bid vector) were defined through a pilot study
179 administered in person to 50 respondents. The bid vector was explored with an open-ended question
180 covering the 15th, 50th, and 85th percentiles; as for the payment vehicle, two methods were tested: a
181 levy on the domestic water bill and an annual donation to a voluntary contribution fund. The majority of
182 respondents preferred the mandatory levy, considered to be fairer because the cost would be evenly
183 distributed among all citizens benefiting from the proposed policy. Monetary contributions would be
184 allocated to a public compensation scheme supporting mountain farmers facing increased irrigation costs
185 under the WFD pricing policy, aiming to ensure the continuation of traditional irrigation practices and
186 the ES they sustain.

187 Table 1 shows the final set of attributes and levels used in the experiment.

189 Table 1. Attributes and levels used in the experiment.

| Attributes | Description | Levels (potential changes within the next 30 years) |
|--|--|--|
| Number of active livestock farms | Maintaining a viable agricultural sector to preserve the economic and social vitality of mountain areas and prevent depopulation | <ul style="list-style-type: none"> - 1,000 (current number of livestock farms)¹ - 800 (20% loss of livestock farms) - 600 (40% loss of livestock farms), non-intervention scenario |
| Number of landslides and debris flows/year | Contributing to the prevention of hydrogeological instability through the maintenance of the existing irrigation network, linked to irrigation management activities | <ul style="list-style-type: none"> - 75 (current number of landslides and debris flows/year)² - 100 (+25 events/year) - 125 (+50 events/year), non-intervention scenario |
| Number of summer tourists/year | Attracting tourists during the summer season, the green landscape enhances the presence of visitors engaged in hiking and outings | <ul style="list-style-type: none"> - 600,000 (current number of summer tourists/year)³ - 480,000 (20% loss) - 360,000 (40% loss), non-intervention scenario |
| Landscape appearance | Preserving the aesthetic value of the landscape through the maintenance of pastoral activities based on currently existing irrigated pastures and grassland |  <p style="text-align: center;">* non-intervention scenario</p> |
| Domestic water bill increase (euro/year) | Willingness to pay for the described level of external benefits | <ul style="list-style-type: none"> - 20 - 50 - 80 - 0 (non-intervention scenario) |

190 ¹ Regional farm archive, year 2020.191 ² Regional archive of hydrogeological disturbance, mean data of the years 2017-2018-2019 (<http://catastodissesti.partout.it/>).192 ³ Regional archive of tourism, number of tourists from June to September, mean data of the years 2017-2018 -2019
193 (https://www.regione.vda.it/asstur/statistiche/default_i.asp).

194

195

196 By repeated choices between bundles of attribute levels (choice sets), respondents evaluated trade-
197 offs among alternatives. In a full factorial design with four attributes at two levels each and one attribute
198 at three levels, the total number of possible profiles is $2^4 \times 3 = 48$. When constructing choice sets composed
199 of two alternatives plus a fixed status quo (non-intervention scenario), the total number of unique
200 unordered choice sets would be $(48 \times 47) / 2 = 1,128$. Given that this number is prohibitively large to present

201 to our sample, the attribute level combinations were generated using an efficient experimental design.
202 We first constructed a D-optimal design under zero priors, assuming no prior information on the sign or
203 magnitude of the parameters¹. After collecting data from the first 50 respondents, we estimated a
204 preliminary model and used the resulting parameter estimates as informative priors. We then re-
205 optimised the design to improve its statistical efficiency, while preserving the same attribute space and
206 level balance. The final experimental design comprised 12 choice tasks, organised into three blocks of
207 four tasks each; respondents were randomly assigned to one of these blocks. Each choice task presented
208 three mutually exclusive irrigation scenarios: two policy alternatives whose attribute levels varied across
209 tasks according to the experimental design, and a *no-intervention* alternative. The *no-intervention*
210 scenario was kept constant across all choice cards.

211

212 *2.2 Questionnaire design and administration*

213 The survey questionnaire was divided into three sections. The introductory section provided a brief
214 overview of the research objectives and information about the study area. In the second section, the
215 hypothetical scenario was presented, along with a description of the ES provided by irrigation of
216 meadows and pastures, followed by the choice experiment. The final section was dedicated to collecting
217 sociodemographic data relevant to individuals' WTP.

218 In line with the experimental design, each of the three blocks featured four choice tasks. Each task
219 presented participants with two alternatives, as well as a *no-intervention* option at zero cost. Figure 1
220 provides an example of one of these choice tasks.

221




222

¹ The design was generated in Stata using the *dcreate* module, which implements a modified Fedorov algorithm (Cook and Nachtsheim, 1980; Zwerina et al., 1996; Carlsson and Martinsson, 2003) to maximize D-efficiency of the design using the covariance matrix of the conditional logit model as input.

223 Figure 1. Example of choice task.

224 The following table shows three alternative scenarios for the Aosta Valley over the next 30 years, after the
225 implementation of the EU water pricing policy. Scenario 1 and Scenario 2 display the benefits of irrigation
226 maintained through citizens monetary contributions, and the resulting increase in domestic water bill per
227 household/year. The third scenario shows the situation without any contribution, at zero costs for the citizens.

228 Which scenario would you prefer?

| | Scenario 1 (in 30 years) | Scenario 2 (in 30 years) | Scenario 3 (in 30 years) |
|--|---|--|---|
| Number of active livestock farms | 1,000 | 800 | 600 |
| Number of landslides and debris flows/year | 100 | 75 | 125 |
| Number of summer tourists | 600,000 | 480,000 | 360,000 |
| Landscape |  |  |  |
| Domestic water bill increase (euro/year per household) | 20 | 80 | 0 |
| Which do you prefer? | <input type="checkbox"/> | <input type="checkbox"/> | <input type="checkbox"/> |

229

230 Respondents who consistently selected the non-intervention scenario were asked to state their reasons,
231 to distinguish between *genuine zeros* and *protest zeros* (Meyerhoff and Liebe, 2006). *Genuine zeros*
232 identify respondents unconcerned about changes projected 30 years ahead or who believe reduced
233 irrigation-related ES would not affect residents' well-being. *Protest zeros* refer to those rejecting aspects
234 of the survey, often perceived as unfair (Jørgensen et al., 2020). In this study, the most common
235 motivations were that existing public funds were deemed sufficient, that farmers already receive
236 adequate EU subsidies, or that respondents preferred voluntary one-off contributions.

237 Data collection was outsourced to an external professional survey agency. The questionnaire was
238 administered online in January 2022 using a CAWI (Computer Aided Web Interviewing) method. The
239 target population consisted of residents of the Aosta Valley aged 18 to 65, with stratified sampling by
240 age and gender applied to reflect the regional population structure. The final dataset consists of 150
241 respondents whose main sociodemographic characteristics are reported in Table 2. Sample data show

242 some discrepancies with regional data on education and income, which should be interpreted with
 243 caution given their imperfect comparability. With these caveats in mind, the sample appears to include
 244 a higher share of lower-income and more highly educated households than the regional population.

245 Table 2. Descriptive statistics of the sample compared with the Aosta Valley population.

| | Sample | Aosta Valley population |
|--------------------------------------|---------------|--------------------------------|
| | Frequency % | Frequency % |
| Gender (female)* | 50.0 | 51.0 |
| Age bracket* | | |
| 19-36 | 28.7 | 28.5 |
| 37-51 | 33.3 | 33.5 |
| 52-66 | 38.0 | 38.0 |
| Education** | | |
| None | - | - |
| Primary school | 0.7 | 3.0 |
| Secondary school | 28.7 | 34.9 |
| High school diploma | 58.0 | 42.4 |
| Academic degree | 12.7 | 19.7 |
| Members of environmental association | 27.3 | N/A |
| Household net income/month (€) | | |
| < 1,200 | 19.1 | N/A |
| 1,200 - 2,000 | 38.2 | N/A |
| 2,000 - 3,000 | 30.9 | N/A |
| > 3,000 | 11.8 | N/A |
| Household yearly net income | | |
| Mean * | N/A | 39,461 € |
| Median * | N/A | 32,900 € |

246 Note: *ISTAT data refer to 2022 **ISTAT data refer to 2020 for individual aged 25-64

247

248 2.3. Method

249 To estimate the utility associated with alternative irrigation scenarios, we adopt a random-parameter
 250 (mixed) logit model in the random utility maximisation tradition (McFadden, 2001; Hensher et al.,

251 2015). The indirect utility that respondent n derives from choosing irrigation policy alternative j in choice
252 situations t specified as:

$$253 \quad U_{njt} = \beta_n' x_{njt} + e_{njt} ,$$

254 Where x_{njt} is a vector of observed attributes describing alternative j (number of livestock farms
255 maintained, number of landslide events avoided, number of summer tourists, landscape quality, and
256 annual increase in the household water bill), β_n is a vector of individual-specific parameters, and e_{njt} is
257 an unobserved error term independently and identically distributed extreme-value type I. Utility is
258 specified as additive in attributes and linear in parameters. Continuous attributes enter the utility function
259 in their natural units, without nonlinear transformations or interaction terms. This implies constant
260 marginal utility over the DCE attribute ranges; therefore, coefficients for active livestock farms,
261 landslides and debris flows, and summer tourists should be interpreted as average marginal effects within
262 the experimentally defined intervals.

263 In the mixed logit specification, preference heterogeneity is captured by allowing the elements of β_n to
264 vary across individuals according to continuous probability distributions. In our baseline model, all non-
265 cost parameters are treated as random and assumed to follow a normal distribution, thereby allowing for
266 continuous preference heterogeneity across individuals, whereas the cost parameter is kept fixed.

$$267 \quad \beta_n = \beta + \eta_n , \eta_n \sim N(0, \Sigma), \quad \text{eq.1}$$

268 Where β is the vector of mean coefficients and Σ is the variance–covariance matrix of the random
269 parameters. Under this specification, the marginal utilities of the non-cost attributes, and therefore the
270 implied willingness-to-pay measures, are allowed to vary across respondents.

271 The model parameters are estimated by simulated maximum likelihood. We use Stata 18 (StataCorp
272 LLC, College Station, TX) and approximate the integrals over the random parameters with 500 Halton
273 draws per respondent. From the estimated coefficients, we derive marginal WTP for each attribute as
274 the ratio between the attribute coefficient and the (negative) cost coefficient, and then compute scenario-
275 specific WTP as appropriate linear combinations of the marginal WTPs.

276
$$WTP_k = -\frac{K}{\text{cost}}, \quad \text{eq. 2}$$

277 For attributes measured on a continuous scale (e.g. number of livestock farms, number of landslides
 278 avoided), WTP_k represents the amount that respondents are willing to pay per year for a one-unit increase
 279 in attribute k . For categorical attributes modelled with dummy variables (e.g. landscape quality), the
 280 corresponding coefficient and WTP measure the change relative to the *non-intervention* scenario.

281 In addition to marginal WTPs, we derive each respondent's total WTP for a policy scenario that
 282 maintains the current provision of irrigation-related benefits over the next 30 years. Let x_0 denote the
 283 vector of attribute levels (including the non-intervention dummy and zero cost) describing this non-
 284 intervention scenario, and x_1 the vector of attribute levels describing the scenario in which the current
 285 provision of irrigation benefits is maintained in 30 years. Using the individual-specific coefficients β_n ,
 286 we compute the respondent-specific compensating surplus (i.e. total WTP) associated with the current
 287 provision of irrigation-related benefits for respondent n as (Mariel et al., 2025):

288
$$WTP_{n,\text{maint}} = -\frac{1}{\text{cost}} \beta'_n (x^1 - x^0), \quad \text{eq.3}$$

289 This expression represents the monetary equivalent of the utility change between the non-intervention
 290 and maintenance scenarios, obtained as the sum of the marginal willingness-to-pay values for each
 291 attribute multiplied by the corresponding attribute change. This yields an empirical distribution of total
 292 WTP across respondents. Aggregate regional values are obtained by multiplying the 25th, 50th (median),
 293 and 75th percentiles of this per-household WTP by the number of households in the Aosta Valley.

294
 295 **3. RESULTS**

296 Nineteen respondents out of 150 (12.6%) chose the non-intervention scenario across all choice sets. All
 297 these zero values were classified as protest zeros, with respondents citing feelings of unfairness about
 298 contributing to irrigation costs, as farmers already receive significant public subsidies. These protest

299 zeros were excluded from the sample, as they do not represent respondents' "true" valuations of the
300 benefits provided by irrigation.

301 Results of the econometric analysis conducted on the 131 valid questionnaires are reported in Table 3.
302 The analysis is based on 524 choice situations (corresponding to 1,572 alternatives).

303 Most of the estimates from the random parameter (mixed) logit analysis were significant ($p < 0.05$),
304 indicating that the attribute had a substantial impact on the respondents' choices. The only attribute that
305 seems to not significantly influence respondents' choices is the number of summer tourists. For the
306 continuous variables, the sign of the estimates indicated a negative (-) or a positive (+) relationship
307 between the respondents' utility and the marginal increase in attribute level. As expected, the
308 respondents' utility increased with an increase in active livestock farms, while it decreased with an
309 increase in landslides and debris flows, as well as with an increase in the domestic water bill. For
310 landscape appearance, respondents derive higher utility from green irrigated meadows compared to
311 degraded landscapes. While the "green landscape" level has a strong and significant positive effect, the
312 "yellow landscape" level is not statistically significant.

313 The WTP values reported in Table 3 correspond to marginal WTP. For continuous attributes, these
314 values represent the WTP for a one-unit change in the attribute; for categorical attributes, they represent
315 the WTP for moving from the reference level (non-intervention/status quo) to the level described in the
316 alternative.

317 Respondents' marginal WTP for each livestock farm maintained is about 0.5 €/year, while the marginal
318 WTP for each avoided landslide is about 3.4 €/year. The marginal WTP to maintain the grazed green
319 meadows instead of the landscape described non-intervention scenario is about 126€/year.

320 The positive and significant coefficient associated with the no-intervention (opt-out) alternative suggests
321 that, all else equal, respondents display a baseline preference for the opt-out alternative. When the
322 intervention scenario preserves multiple ES at their current level, their combined utility outweighs the
323 baseline preference for the opt-out alternative.

324

325 Table 3. Mixed logit estimates

| Attribute | Coef. | Std.err | p-value | Marginal WTP (€/year) |
|--|-----------|---------|---------|--------------------------|
| Number of active livestock farms | 0.009*** | 0.00 | <0.001 | 0.50 |
| Number of landslides and debris flows/year | -0.062*** | 0.01 | <0.001 | -3.38 |
| Number of summer tourists (thousands)/year | -0.003 | 0.00 | 0.15 | -0.15 |
| Landscape appearance (green landscape) | 2.312*** | 0.39 | <0.001 | 126.34 |
| Landscape appearance (yellow landscape) | 0.011 | 0.01 | 0.366 | 0.61 |
| Domestic water bill increase (euro/year) | -0.018** | 0.00 | 0.019 | |
| No- intervention (opt-out) | 3.184*** | 1.14 | 0.006 | 173.99 |
| σ (Number of active livestock farms) | 0.007*** | 0.002 | 0.004 | |
| σ (Number of landslides and debris flows) | 0.058*** | 0.009 | <0.001 | |
| σ (Number of summer tourists) | 0.016*** | 0.004 | <0.001 | |
| σ (Landscape appearance - green) | 2.388*** | 0.436 | <0.001 | |
| σ (Landscape appearance - yellow) | 0.261 | 0.287 | 0.363 | |

326 Note: σ indicates the standard deviation of the random parameters distribution. Random parameters are assumed to be normally
327 distributed. (*p < 0.10; **p < 0.05; *** p < 0.01); Log-likelihood = -448.6; McFadden's pseudo-R² = 0.32; AIC =933.2; BIC =1032.2;
328 Observations (choice cards analyzed) = 524; Respondents = 131.

329

330 Beyond marginal values, we also calculated the total WTP associated with a scenario that preserves the
331 current provision of irrigation-related benefits over the next 30 years. Specifically, the WTP values were:
332 €126.34/year for maintaining grazed green meadows, €200/year to preserve the current number of 400
333 livestock farms, €169/year to prevent an increase of 50 landslides, -€36/year for preserving the current
334 number of summer tourists, and €173.99/year for the opt-out scenario. The calculations are specified in
335 Appendix A.

336 The median value calculated per household is about 633.5 €/year (Table 4). Based on this value and the
337 number of households living in the Region, we estimated a total benefit of almost 36 million €/year for
338 the regional population.

339

340 Table 4. Willingness to pay for the best scenario (€/year).

| | WTP (25% percentile) | WTP (50% percentile) | WTP (75% percentile) |
|--|----------------------|----------------------|----------------------|
| Total value per household (€) | 564.26 | 633.35 | 705.18 |
| Aggregated regional value (€) (56,562 households ¹) | 30,897,558 | 35,823,542 | 39,886,391 |

341 ¹ <https://demo.istat.it/previsionifamiliari/index.php?lingua=ita>

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Finally, following Weituschat et al., (2023), the heterogeneity of attribute preferences among respondents was analysed, testing the relationship between the marginal WTPs obtained from the mixed logit model and the socioeconomic characteristics of the sample. Table 5 shows the results of a seemingly unrelated regression model, where the individual marginal WTPs for the four attributes (dependent variables) are regressed on the individual socioeconomic characteristics (explanatory variables). The sign of the coefficients indicates whether there is a direct or inverse relationship between attributes' marginal WTPs and either the marginal increase of continuous explanatory variables (age, educational qualification, income bracket) or, in the case of dummy variables, the belonging to a specific category (female, membership in an environmental association).

According to the results, gender does not seem to influence the respondents' WTP, as the estimated coefficients are not significant for any of the considered attributes. Younger respondents place a higher value on the presence of summer tourists than older respondents. The opposite holds true for the benefits from landslides reduction, livestock farms maintenance and agricultural landscape maintenance, which seem to be more appreciated as the age of respondents increases. The value placed on landslides reduction and landscape maintenance is lower for higher educated individuals. As the income bracket increases, the WTP for all the ES increases, except for the number of livestock farms. Members of environmental associations show a lower-than-average WTP for both landslide reduction and agricultural landscape maintenance.

Furthermore, it should be noted that the constants in Table 5 are large and statistically significant for all attributes, indicating that a substantial share of the heterogeneity in marginal WTPs is not captured by the socioeconomic variables included in the model. This suggests that factors other than the observed demographics, play a major role in shaping respondents' valuations.

365

366 Table 5. Seemingly unrelated regression results.

| | Active livestock farms | Landslides and debris flow | Summer tourists | Landscape appearance |
|---------------------------|------------------------|----------------------------|-----------------|----------------------|
| Age (years) | 0.001* | -0.013** | -0.004*** | 0.448*** |
| Gender (male) | 0.01 | -0.081 | -0.029 | -0.841 |
| Education (level) | -0.001 | 0.131** | 0.021 | -7.995*** |
| Income (bracket) | -0.027*** | -0.194*** | 0.024* | 20.275*** |
| Member of env. ass. (yes) | 0.013 | 0.192** | 0.008 | -15.261*** |
| Constant | 0.536** | -3.344*** | -0.171* | 114.521*** |

367 Note: (*p < 0.10; **p < 0.05; *** p < 0.01)

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4. DISCUSSION

371 Our results provide insights into the perceptions of the citizens of the Aosta Valley regarding the ES
 372 provided by the irrigation of meadows and pastures.

373 Although European citizens generally support public aid to farmers (Mata & Dos-Santos, 2024; Schüler
 374 & Noack, 2019; European Commission, 2016), more than one in ten respondents in our survey protested
 375 against further support for local farmers. This may be partly explained by the Aosta Valley's special
 376 statute status, where substantial regional funds complement EU support. The same context may also
 377 underlie the baseline preference for the no-intervention option, suggesting reluctance to support
 378 additional financial contributions despite recognising the value of the ES provided.

379 Yet, citizens with a positive WTP expressed a significant appreciation for the external benefits of
 380 irrigation, probably due to their strong identification with the distinctive features that define the region
 381 (Novelli et al., 2022).

382 This finding aligns with evidence from other EU regions, where agriculture and forestry are broadly
 383 perceived as providers of public goods and payments for ES are considered a relevant policy response
 384 (Targetti et al., 2022). Positive WTP for agri-environmental, landscape and irrigation-related ES has

385 similarly been documented across other EU contexts, including Flanders, Ireland, the Po Valley and
386 Lombardy (Fockaert et al., 2023; Howley et al., 2012; Blasi et al., 2023; Ruberto et al., 2022).

387 Tourism is the only non-significant attribute, likely reflecting heterogeneous preferences for visitors
388 across socioeconomic groups (Muresan et al., 2021; Şorcaru et al., 2022; Brida et al., 2011). In the Aosta
389 Valley context, this heterogeneity is mainly explained by a generational tension: older respondents are
390 possibly more concerned about cultural disruption, while younger ones focus on the expected economic
391 benefits of tourism.

392 Respondents valued livestock farm maintenance, confirming that citizens recognise the ES provided by
393 traditional agricultural practices, including food security, water regulation and cultural heritage
394 (Faccioni et al., 2019; García-Llorente et al., 2012; Tempesta & Vecchiato 2018). WTP for this attribute
395 is negatively correlated with family income, suggesting that livestock maintenance may be perceived as
396 an 'inferior good' (Ebert, 2003): lower-income respondents may fear that fewer farms would reduce
397 market competition and raise prices, or value the direct economic benefits of livestock activities, such
398 as employment and access to local products.

399 Respondents showed considerable concerns about landslide risk, reflected in the high WTP estimated
400 for this attribute, consistent with findings from similar studies (Alcon et al., 2020; Owuor et al., 2019).
401 WTP for landslide reduction is higher among younger respondents and members of environmental
402 organisations, likely reflecting greater awareness of current environmental risks, particularly in
403 mountainous areas considered most vulnerable to climate change (Palomo-Campesino et al., 2018).

404 The value placed on maintaining the traditional agricultural landscape was significant and positive,
405 consistent with similar studies (Targetti et al. 2020; Faccioni et al., 2019; Notaro et al. 2019; Novikova
406 et al., 2019; Dupras et al., 2018; Tempesta & Vecchiato 2018). Interestingly, WTP decreases with higher
407 education and membership of environmental organisations, possibly reflecting a preference for
408 wilderness and renaturation over traditional heritage landscapes, and greater sensitivity to water scarcity.

409 Direct comparisons of WTP estimates from stated-preference methods are limited by differences in
410 context, scale, beneficiaries, and valuation design. Against this backdrop, our estimated median WTP of
411 about 630 €/year per household for the current provision scenario lies in the upper range of comparable

412 European studies. Ruberto et al. (2022) report WTP values of approximately 30-40 €/month for
413 preserving ES from traditional irrigation systems in Lombardy (corresponding to roughly 360-480 €/year
414 per household). Our estimate reflects the aggregation of multiple benefits within a single policy
415 scenario, which partly explains its higher magnitude. A comparable DCE by Faccioni et al. (2019),
416 estimated a total economic value of 159.30 €/year per person for four ES (landscape, biodiversity, water
417 quality, and quality local products) in the Alpine agroecosystem of Trento Province, broadly comparable
418 to our figure once differences in aggregation level are accounted for. It is also worth noting that the
419 higher total WTP may also reflect both the 30-year time horizon and the Alpine context, characterized
420 by strong landscape identity and sensitivity to hydrogeological risks, which likely reinforce stated WTP
421 compared to shorter-term or less risk-exposed settings.

422

423 5. CONCLUSION

424 The WFD requires water pricing across all sectors, including agriculture, where farmers should cover
425 the financial, environmental and resource costs of irrigation. In Italy, the Aosta Valley has been criticised
426 for lacking a compliant pricing system, raising concerns about affordability for local livestock farmers
427 (Borsotto et al., 2021). However, irrigation in mountain areas also generates ecosystem services and
428 environmental benefits that create a socially derived demand for water and justify internalising these
429 amenities through water allocation policies (Thiene & Tsur, 2013). Monetising them is therefore
430 necessary for their recognition at EU level and for comparison with irrigation costs.

431 The total annual benefit of irrigation for regional residents was estimated at around €36 million, a
432 substantial figure given that only four ES were valued. Irrigation, in fact, generate a broader range of
433 benefits not captured here, including water quality maintenance, firefighting infrastructure, and
434 professional role creation (Novelli et al., 2022).

435 Despite differences in context, scale and methods, the findings accord with European evidence of
436 positive WTP for landscape preservation, hydrogeological risk reduction and traditional mountain

437 farming (Faccioni et al., 2019; Bernués et al., 2019; Novikova et al., 2017; Dupras et al., 2018). Tourism
438 was the only exception, with conflicting preferences that require careful management.

439 These findings should nonetheless be interpreted with caution. Partial interdependence among attributes
440 - which reflects real-world irrigation dynamics -means that additive WTP may overstate the value of the
441 whole package, as respondents may have valued certain combinations as a bundle rather than as
442 independent goods. The small sample (131 after excluding protest zeros), though comparable to other
443 regional DCE studies (Blasi et al., 2023; García-Llorente et al., 2012), limits statistical power and
444 generalisability.

445 These results are therefore indicative and context-specific, yet they offer a contribution to the broader
446 debate on balancing agricultural practices with sustainable resource use, particularly in informing the
447 design of socially sustainable agri-environmental policies.

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Appendix A. Calculation of individual willingness to pay for the policy scenario

| Attribute | Marginal WTP (€/year) | Δ (policy scenario – baseline) | Total (€/year) (Marginal WTP \times Δ) |
|--|--------------------------|--|---|
| Number of active livestock farms | 0.50 | 400 | 200 |
| Number of landslides and debris flows/year | -3.38 | -50 | 169 |
| Number of summer tourists (thousands)/year | -0.15 | 240 | -36 |
| Landscape appearance (green landscape) | 126.34 | | 126.36 |
| No- intervention (opt-out) | 173.99 | | 173.99 |
| Total | | | 633.35 |

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Just Accepted